# T-61.5040 Oppivat mallit ja menetelmät <br> T-61.5040 Learning Models and Methods <br> Pajunen, Viitaniemi 

Exercises 8, 16.3.2007

## Problem 1.

Assume that the true posterior is $p(\theta \mid y)=N\left(\theta \mid 0, \sigma^{2} I\right)$, where $\theta \in \mathbb{R}^{1000}$. You try to perform rejection sampling using a proposal density $g(\theta)=N\left(\theta \mid 0, \sigma_{g}^{2} I\right)$. Suppose that your proposal density is close to the true posterior so that $\sigma_{g}=1.1 \sigma$. Compute approximately how many samples are rejected.

## Problem 2.(demonstration)

The Metropolis algorithm simulates a posterior by starting from a value $\theta^{0}$. Then the algorithm repeats a step $n$ which produces value $\theta^{n}$ give $\theta^{n-1}, n=1,2,3, \ldots$ In the $n$th step of the algorithm a new value $\theta^{*}$ is picked from a jumping distribution $J\left(\theta^{*} \mid \theta^{n-1}\right)$. The new value is accepted $\left(\theta^{n}=\theta^{*}\right)$ with a probability $p_{r}=\min \{1, r\}$ where $r=\frac{p\left(\theta^{*} \mid y\right)}{p\left(\theta^{n-1} \mid y\right)}$. If it is not accepted, the next sample is $\theta^{n}=\theta^{n-1}$.

Assume that the Markov chain defined by the Metropolis algorithm has a unique stationary distribution. Show that this distribution is the posterior $p(\theta \mid y)$ (assume that the jumping distribution is symmetric).

What if the posterior is $1 / 2 p_{1}(\theta \mid y)+1 / 2 p_{2}(\theta \mid y)$, where $p_{1}$ and $p_{2}$ are uniform distributions over $[0,1]$ and $[2,3]$, respectively? Can you think of a jumping distribution that prevents a unique stationary distribution?

## Problem 3.

Observe data $y_{1}, \ldots, y_{n}$ from a Normal distribution $N\left(\mu, \sigma^{2}\right)$. Assume both $\mu$ and $\sigma^{2}$ are unknown. Choose the priors as $p(\mu \mid \sigma)=N\left(\mu \mid \mu_{0}, \sigma_{0}^{2}\right)$ and $p\left(\sigma^{2}\right)=I G\left(\sigma^{2} \mid a, b\right)$.
i) Formulate the Gibbs Sampler for the unknowns $\mu$ and $\sigma^{2}$.
ii) Describe how you estimate the posterior mean of $\mu$ using the simulated posterior. Recall that Monte Carlo approximation is $\mathrm{E}\left(h\left(\mu, \sigma^{2}\right) \mid y\right) \approx \frac{1}{N} \sum_{i} h\left(\mu_{i}, \sigma_{i}^{2}\right)$ where $\left(\mu_{i}, \sigma_{i}^{2}\right)$ is the i:th simulated posterior sample.

Hint: inverse-gamma distribution is $I G(z \mid a, b) \propto z^{-(a+1)} \exp (-b / z)$. It is a conjugate prior for $\sigma^{2}$ when the model is $N\left(\mu, \sigma^{2}\right)$ where $\mu$ is known. Writing $y=\frac{1}{n} \sum_{i}\left(y_{i}-\mu\right)^{2}$, the posterior for $\sigma^{2}$ is

$$
p\left(\sigma^{2} \mid D\right)=I G\left(\sigma^{2} \left\lvert\, \frac{n}{2}+a\right., \frac{1}{2}(2 b+n y)\right) .
$$

## Problem 4.

You have observed independent samples $y_{1}, \ldots, y_{n}$ from a distribution

$$
p(y \mid a, b)=b \exp (-b(y-a)), \text { when } y \geq a \text { and } p(y \mid a, b)=0, \text { when } y<a .
$$

The parameters $a$ and $b$ are nonnegative.
i) Choose an uninformative prior $p(a, b) \propto b^{-1}$ and compute the unnormalized posterior of $a$ and $b$. Which two scalar functions of data $y_{1}, \ldots, y_{n}$ determine the posterior?
ii) Write a Gibbs Sampler for the posterior. Hint: Gamma distribution for $x$ is $\operatorname{Gamma}(\theta \mid c, d) \propto x^{c-1} \exp (-d x)$.
Another hint: you don't have to simulate $a$ if it seems difficult. You might want to simulate $\exp (a b n)$, since you can solve $a$ from these values when $b$ is known.

